

Multidimensional Buffered Probability of Exceedance: Dual Representations and Multivariate Stochastic Dominance

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Outline

- ① Conservative distribution approximations
- ② Buffered probability of exceedance (bPOE)
- ③ Multidimensional bPOE (M-bPOE)
- ④ (Multivariate) conservative distribution approximations

Problem Description

- Sample distribution of deviations from a desired value.
- Robust approximation is needed (preferences: continuous/grid, closed form).
- Distribution is one of the many in a system. Deviations from central values add up. Risk for the system.
- Conservative approximation needed: risks of high deviations must not be underestimated.

Optimization Problem

$$\begin{aligned} \max_Y \quad & \mathcal{H}(Y) \\ \text{s.t.} \quad & X \leq'_2 Y, -X \leq'_2 -Y, \\ & \sigma^2(Y) \leq \sigma^2(X) + \Delta\sigma^2. \end{aligned}$$

- ① Why Second-Order Stochastic Dominance?

Risk averse decision maker

- ② Why Entropy Maximization?

Smooth and robust solution

- ③ Why Variance Constraint?

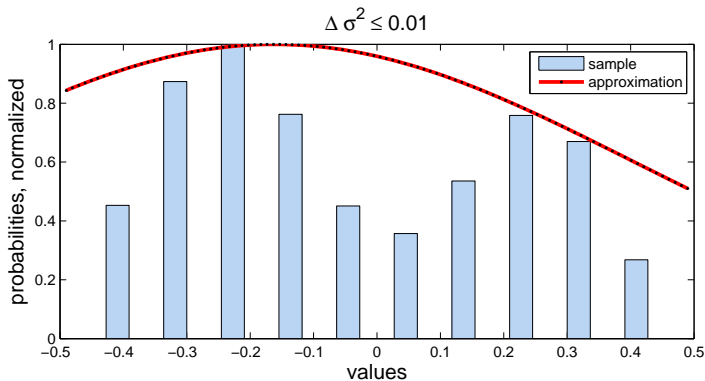
Bounded feasible set

Method properties

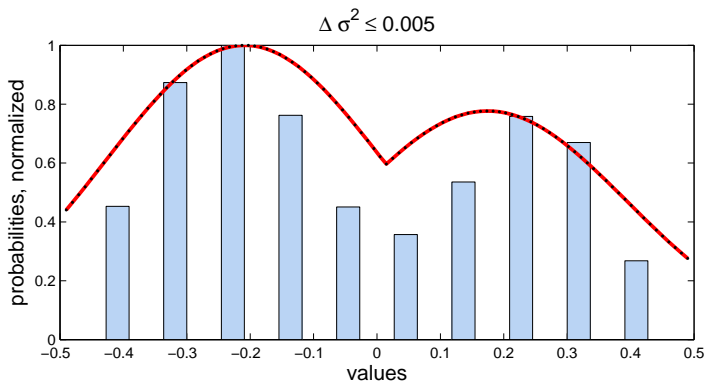
- Discussed dominance constraint guarantees “fatter” right and left tails compared to sample X
- When $\Delta\sigma^2 \rightarrow 0$, then optimal solution converges in distribution to the sample distribution
- Optimal solution is a maximum-of-Gaussians distribution

$$P(Y) = \max_{i=1,\dots,m} C_i \exp\{-(Y - \mu_i)^2 / 2\sigma^2\}$$

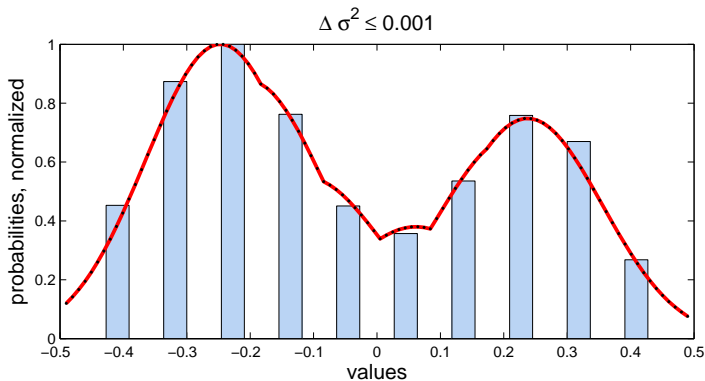
Convergence Example 1: $\Delta\sigma^2 \leq 0.01$



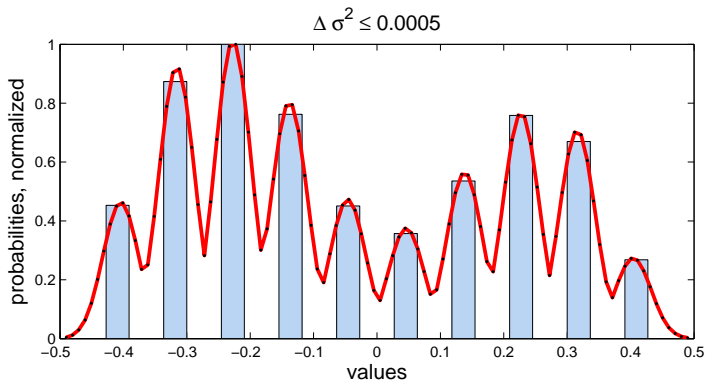
Convergence Example 2: $\Delta\sigma^2 \leq 0.005$



Convergence Example 3: $\Delta\sigma^2 \leq 0.001$

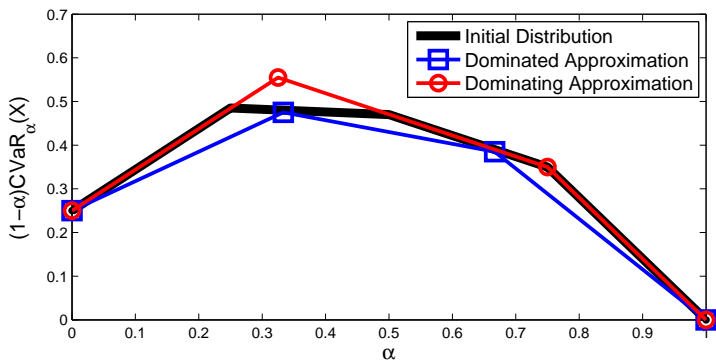


Convergence Example 4: $\Delta\sigma^2 \leq 0.0005$



Finite number of superquantile constraints

For discrete distribution the scaled superquantile is
$$(1 - \alpha_i) \bar{Q}_{\alpha_i}(X) = \sum_{j=i+1}^m p_j x_j, \quad \alpha_0 = 0, \quad \alpha_i = \sum_{j=1}^i p_j$$



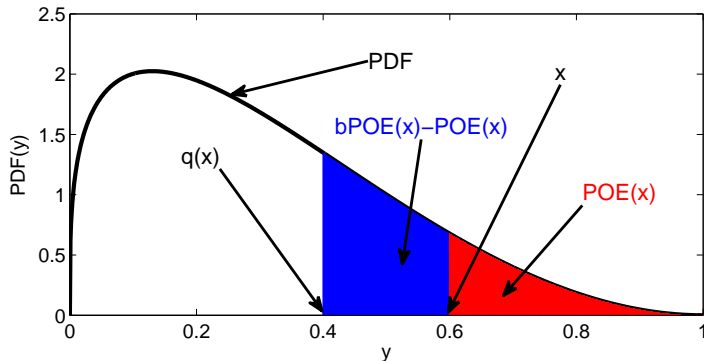
$\bar{Q}_\alpha(Y) \geq \bar{Q}_\alpha(X)$ for $\alpha = \alpha_i$ is sufficient for SOSD

bPOE explanation (continuous case)

$P(X > x)$: probability of exceedance

$P(X > q(x))$: *buffered* probability of exceedance

$q(x) : E[X|X > q(x)] = x$, $x - q(x)$: “buffer”



Buffered Probability of Exceedance

$$P(X \geq x) \equiv p_x(X) \leq \bar{p}_x(X)$$

$\bar{p}_x(X)$ is the only smallest **quasi-convex** and **law-invariant** upper bound for $p_x(X)$

$$\bar{p}_x(X) = \inf_{a \geq 0} E[a(X - x) + 1]^+$$

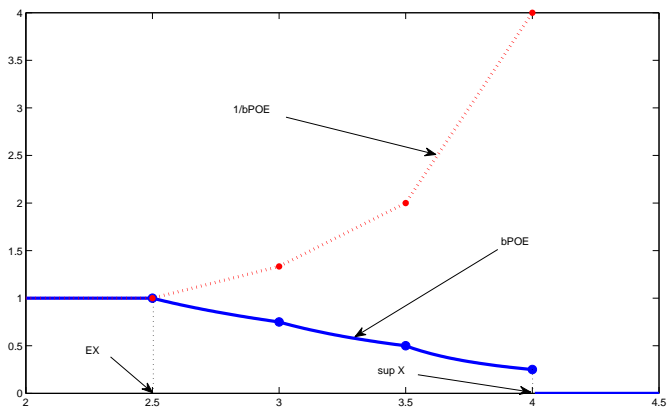
where $[x]^+ = \max\{0, x\}$

$$\bar{p}_x(X) = 1 \text{ for } x < EX \quad \bar{p}_x(X) = 0 \text{ for } x > \sup X$$

$\bar{p}_x(X)$ is decreasing and continuous for $x \in [EX, \sup X)$

bPOE Properties w.r.t. Parameter x

$1/\bar{p}_x(X)$ is a convex nondecreasing function of x and piecewise-linear if X is discretely distributed



Primal and Dual formulations for bPOE

Primal

$$\bar{p}_x(X) = \min_{a \geq 0} E[a(X - x) + 1]^+$$

Dual

$$\begin{aligned} \bar{p}_x(X) = \max \quad & EW \\ \text{s.t.} \quad & EXW \geq xEW \\ & 0 \leq W \leq 1 \end{aligned}$$

Dual Multivariate

$$\begin{aligned} \bar{p}_x(\mathbf{X}) = \max \quad & EW \\ \text{s.t.} \quad & EX_i W \geq x_i EW, \quad i = 1, \dots, d \\ & 0 \leq W \leq 1 \end{aligned}$$

Primal Multivariate

$$\bar{p}_x(\mathbf{X}) = \min_{\mathbf{a} \geq 0} E[\mathbf{a}^T (\mathbf{X} - \mathbf{x}) + 1]^+$$

Suppose that instead of random variable X there is a random vector $\mathbf{X} = (X_1, \dots, X_n)$.

If the law g for aggregating components of \mathbf{X} into a value $g(\mathbf{X})$, determining undesired events, is known, then the problem is reduced to bPOE framework.

What if there is no known law? Let us quantify risks as follows:

$$F_{-\mathbf{X}}(-\mathbf{x}) = P(X_1 \geq x_1 \text{ and } \dots \text{ and } X_n \geq x_n) \equiv p_x(\mathbf{X}) \leq \bar{p}_x(\mathbf{X})$$

Lift Zonoid (by Koshevoy & Mosler)

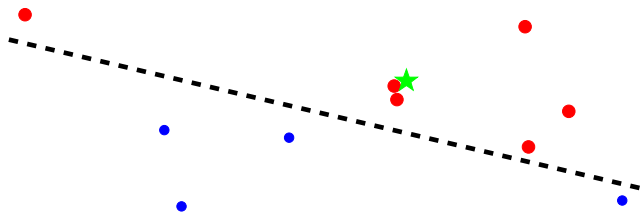
$$Z(\mathbf{X}) = \{(p, y) \in \mathbb{R}^{d+1} | p = EW, y_i = EWX_i, 0 \leq W \leq 1\}$$

$Z(\mathbf{X})$ uniquely determines distribution of \mathbf{X} ;

$Z(\mathbf{X})$ is a convex set; $(0, \mathbf{0}) \in Z(\mathbf{X})$; $(1, E\mathbf{X}) \in Z(\mathbf{X})$;

if $E|\mathbf{X}| < \infty$, then $Z(\mathbf{X})$ is compact;

if \mathbf{X} is discretely distributed, then $Z(\mathbf{X})$ is polyhedral
extreme points \leftarrow expectations within linear cut-offs



Lift Zonoid Transformation

Homography $(p, \mathbf{u}) \rightarrow (1/p, \mathbf{u}/p)$ applied to $Z(\mathbf{X})$

line \rightarrow line & continuous in $\mathbb{R}_{++} \times \mathbb{R}^d \Rightarrow$ preserves convexity

Transforms $Z(\mathbf{X})$ into $\text{epi}(1/f_{\mathbf{X}}(\mathbf{x}))$, where

$$\begin{aligned} f_{\mathbf{X}}(\mathbf{x}) &= \sup_W EW &= \inf_{\mathbf{a}} E[\mathbf{a}^T(\mathbf{X} - \mathbf{x}) + 1]^+ \\ \text{s.t. } &EW\mathbf{X} = \mathbf{x}EW \\ &0 \leq W \leq 1 \end{aligned}$$

$1/f_{\mathbf{X}}(\mathbf{x})$ is convex, piecewise-linear, with knots of type

$$\bar{\mathbf{x}}(\lambda, b) = \sum_{\lambda^T \mathbf{x}^i \geq b} \mathbf{x}^i p^i / \sum_{\lambda^T \mathbf{x}^i \geq b} p^i, \quad \mathcal{X}_{\mathbf{X}} := \bigcup_{\lambda, b} \bar{\mathbf{x}}(\lambda, b)$$

Stochastic Dominance Equivalence

Zonoid Dominance (Mosler & Koshevoy):

$$\mathbf{X} \leq_Z \mathbf{Y} \Leftrightarrow Z(\mathbf{X}) \subseteq Z(\mathbf{Y})$$

Hence, $\mathbf{X} \leq_Z \mathbf{Y} \Leftrightarrow f_{\mathbf{X}}(z) \leq f_{\mathbf{Y}}(z)$ for $z \in \mathcal{X}_{\mathbf{X}}$

$\mathbf{X} \leq_Z \mathbf{Y}$ constraint as a finite number of linear constraints!

Linear Second Order Dominance (Dentcheva & Ruszczyński):

$$\mathbf{X} \leq_2^{\text{lin}} \mathbf{Y} \Leftrightarrow \mu^T \mathbf{X} \leq_2 \mu^T \mathbf{Y} \text{ for all } \mu \geq 0$$

$$\mathbf{X} \leq_2^{\text{lin}} \mathbf{Y} \Leftrightarrow \bar{p}_z(\mathbf{X}) \leq \bar{p}_z(\mathbf{Y}) \text{ for } z \in \mathcal{X}_{\mathbf{X}}^+ := \cup_{\mu \geq 0, b} \bar{x}(\mu, b)$$

$\mathbf{X} \leq_2^{\text{lin}} \mathbf{Y}$ constraint as a finite number of linear constraints!

$\mathbf{X} \leq_Z \mathbf{Y}$ and $\mathbf{X} \leq_2^{\text{lin}} \mathbf{Y}$ are closely related as $f_{\mathbf{X}}(\mathbf{x})$ and $\bar{p}_{\mathbf{x}}(\mathbf{X})$:

$$f_{\mathbf{X}}(\mathbf{x}) = \inf_{\mathbf{a} \in \mathbb{R}^d} E[\mathbf{a}^T (\mathbf{X} - \mathbf{x}) + 1]^+ \quad \bar{p}_{\mathbf{x}}(\mathbf{X}) = \inf_{\mathbf{a} \succeq 0} E[\mathbf{a}^T (\mathbf{X} - \mathbf{x}) + 1]^+$$

Optimization problem formulation

Take $\gamma > 1$ (analogue of $\Delta\sigma^2$ in 1-dimensional version)

$$\begin{array}{ll}\max_{\mathbf{Y}} & \mathcal{H}(\mathbf{Y}) \\ \text{s.t.} & \mathbf{X} \leq_z \mathbf{Y} \quad (\Leftrightarrow -\mathbf{X} \leq_z -\mathbf{Y}) \\ & \sigma^2(\mu^T \mathbf{Y}) \leq \gamma \cdot \sigma^2(\mu^T \mathbf{X}) \quad \text{for all } \mu\end{array}$$

Dominance constraint \Leftrightarrow for all $\bar{\mathbf{x}}^i \in \mathcal{X}_{\mathbf{X}}$:

$$E_{V^i} \mathbf{1} \geq f_{\mathbf{X}}(\bar{\mathbf{x}}^i)$$

$$E_{V^i} \mathbf{Y} = \bar{\mathbf{x}}^i E_{V^i} \mathbf{1}$$

$$0 \leq V^i \leq P$$

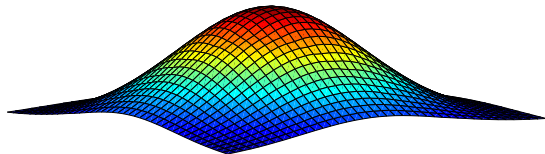
Variance constraint $\Leftrightarrow (\Lambda \succ 0 \Leftrightarrow \Lambda - \text{PSD})$

$$E_P \mathbf{Y} \mathbf{Y}^T - \gamma E \mathbf{X} \mathbf{X}^T + (\gamma - 1) E \mathbf{X} E \mathbf{X}^T \succ 0$$

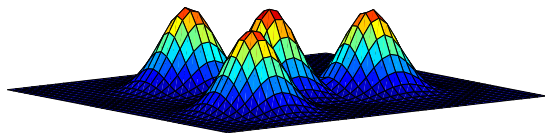
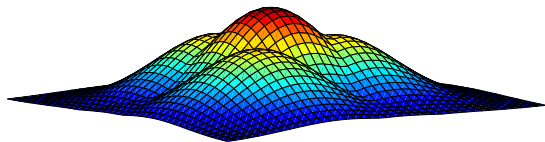
Maximum of Gaussians form of optimal solution

With $\gamma \rightarrow 1$, optimal solution \rightarrow_F sample distribution
Optimal solution is a weighted maximum of Gaussian functions: $\Lambda \succ 0$, $\mu_\alpha \geq 0$

$$P(Y) = \exp \left\{ -\mathbf{Y}^T \Lambda \mathbf{Y} - \lambda_0^T \mathbf{Y} - \lambda_1 + \sum_i [\mu_\alpha^i + (\mathbf{Y} - \bar{\mathbf{x}}^i)^T \lambda_x^i]^+ \right\}$$



Convergence in distribution



Thank you!